Bayesian Estimation of altimeter echo parameters

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In the current retracking procedures, waveforms are processed independently from the previous ones by comparing the measured altimetric waveform with a return power model (Hayne model) according to least square estimators.

Page 2



However, consecutive altimetric waveforms are representative of continuous ocean surfaces and **it seems promising** to account from previous WFs when estimating the ocean parameters

- A first step has been done by JPL (E.Rodriguez) when processing the 10 Wfs of the 1s WF-packet at the same time (Topex retracking in RGDR) and solving for 10 ranges but only one SWH and one sigma0 per second
- We have been investigating a **Bayesian approach** combining information coming from the data (contained in the likelihood) and prior knowledge regarding unknown parameters (contained in the parameter prior distribution)



The altimetric signal is modeled using a Brown model

$$y_k = s_k \cdot e_k$$

k=0,...,K-1 ; K nb of WF samples

Page 3

 \mathcal{C}_k is a multiplicative speckle noise (L: number of integrated pulses) $\mathcal{C}_k \sim G(L,L)$

The likelihood of the estimation vector y (the waveform) is :

$$f(\boldsymbol{y}|\boldsymbol{\theta}) = \left[\frac{L^L}{\Gamma(L)}\right]^K \exp\left(-L\sum_{k=1}^K \frac{y_k}{s_k}\right) \left(\prod_{k=1}^K y_k\right)^{L-1} \left(\prod_{k=1}^K s_k\right)^{-L}$$

where θ is the unknown parameter vector (τ , SWH,Pu,...) y is the observed waveform $y=(y_0,...,y_{K-1})$ with independant noise samples $f(y/\theta)$ is the product of K Gamma probability density functions

The Maximum likelihood Estimator (MLE) of θ is : $\theta_{MV} = \arg \max_{\theta} f(y|\theta)$

- → No analytical expression for the Maximum Likelihood Estimator of θ
- → We currently use a quasi Newton method (approximate solution)

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The posterior distribution of θ is defined as :

 A priori » function also called prior
 distribution

Page 4

 $f(\boldsymbol{\theta}|\boldsymbol{y}) = \frac{f(\boldsymbol{y}|\boldsymbol{\theta})f(\boldsymbol{\theta})}{f(\boldsymbol{y})} \propto f(\boldsymbol{y}|\boldsymbol{\theta})f(\boldsymbol{\theta})$ Likelihood function = information coming from the data

where y is the observed waveform and θ is the unknown parameter vector (τ , SWH,Pu, ...)

Two possibilities to achieve the estimation of θ

- the minimum mean square error (MMSE) (mean of the posterior distribution)
- the maximum a posteriori (MAP) (method investigated some years ago by E.Rodriguez on Topex data)

 $\widehat{\boldsymbol{\theta}}_{MMSE} = \mathbf{E}[\boldsymbol{\theta}|\boldsymbol{y}], \\
\widehat{\boldsymbol{\theta}}_{MAP} = \arg \max_{\boldsymbol{\theta}} f(\boldsymbol{y}|\boldsymbol{\theta}) f(\boldsymbol{\theta})$



Prior distributions summarize the available information regarding the unknown parameters. Different scenarii have been investigated

- Uniform distributions on appropriate intervals,
- Priors computed from parameter estimates along a cycle (fitted by splines),
- Time-varying prior distributions (dynamic distribution=Gaussian distribution centered on the previous estimation).





$$f(P_u, \tau, \text{swh}|\mathbf{y}) \propto \exp\left[-L\sum_{k=1}^{K} \frac{y_k}{s_k}\right] \frac{\left(\prod_{k=1}^{K} y_k\right)^{L-1}}{\left(\prod_{k=1}^{K} s_k\right)^{L}} f(P_u)f(\tau)f(\text{swh})$$

Page 6

- Too complex to derive analytical expression for the MAP and MMSE estimates of θ
- We used Markov Chain Monte Carlo (MCMC) methods which consists in building a Markov Chain $\theta(t)$ to generate samples distributed according to the a-posteriori distribution

(more details in Severini, Mailhes, Thibaut and Tourneret, IGARSS proceedings, Boston, 2008)

• A hybrid "Metropolis-within-Gibbs" algorithm has been used to generate candidates according a given pdf with an accept/reject procedure.

$$\widehat{\theta}_{\text{MMSE}} \simeq \frac{1}{T} \sum_{t=1}^{T} \boldsymbol{\theta}^{(t)}$$



With uniform priors



Page 7

Δ : pseudo MLE, **o** : Bayesian estimator

- Pseudo MLE and Bayes algorithms perform similarly for the parameters Pu and τ
- Better performance for SWH using the Bayesian method.



With dynamic priors



Page 8

Δ : pseudo MLE, **o** : Bayesian estimator

Prior distributions at time instant t are defined as Gaussian densities $\mathcal{M}(m; \sigma^2)$, with $m = \theta^{(t-1)}$ and an appropriate variance σ^2 ,

Performance

- similar for the estimation of Pu
- Gain for τ
- Gain for SWH

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(Jason-2 - tracks 1 to 20 from cycle 16)

Results on SLA

SLA Power Spectrum (on 20Hz estimates)



Results on real data

(Jason-2 - tracks 1 to 20 from cycle 16)

SWH @20-Hz, Ku-Band, Cycle 016 T022

Results on SWH





Results on real data

Results on sigma0

(Jason-2 - tracks 1 to 20 from cycle 16)

- Products
- _____ σ0 from MLE3
- --- σ 0 from Bayesian with MLE4 priors
- **σ0 from Bayesian with filtered SWH priors**



> SIGMA-0 @20-Hz, Ku band, C016 T001-T020, J2 (×10⁴) Nbr = 701358 StdDev = 1.406 products Mov = 13.75 Rtk MLE3 Nbr = 701901 Mov = 13.55 StdDev = 1.351Bavesian (MLE4 swh) Nbr = 701876 StdDev = 1.351 Moy = 13.55 Bayesian (filtered swh Nbr = 699889 Moy = 13.55 StdDev = 1.346 12 14 18 σ0(dB) - Ku-band

Page 11

• Strong noise reduction on this parameter

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- Very promising results
- The method gives the a-posteriori distribution and the confidence interval of the estimations

- However, time consuming method
- This method could be used localy or for very precise estimations (bathymetry) or high rate mean profiles (20Hz)
- Results could be compared with the two passes method developed by NOAA (W.Smith) on typical situations
- Simulations have been done on Median tracking cycles : improvements expected on Diode/DEM cycles. To Be Evaluated ...

